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THE RATE OF CONVERGENCE OF THE MEAN LENGTH OF THE LONGEST COMMON SUBSEQUENCE¹

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Given two i.i.d. sequences of n letters from a finite alphabet, one can consider the length L_n of the longest sequence which is a subsequence of both the given sequences. It is known that EL_n grows like γn for some $\gamma \in [0, 1]$. Here it is shown that $\gamma n \geq EL_n \geq \gamma n - C(n \log n)^{1/2}$ for an explicit numerical constant C which does not depend on the distribution of the letters. In simulations with n = 100,000, EL_n/n can be determined from k such trials with 95% confidence to within $0.0055/\sqrt{k}$, and the results here show that γ can then be determined with 95% confidence to within $0.0225 + 0.0055/\sqrt{k}$, for an arbitrary letter distribution.

1. Introduction. Given a finite alphabet A and two sequences x_1, \ldots, x_n and y_1, \ldots, y_n in A, there is said to be a common subsequence of length k if for some $1 \le i_1 < \cdots < i_k \le n$ and $1 \le j_1 < \cdots < j_k \le n$, we have $x_{i_m} = y_{j_m}$ for all $1 \le m \le k$. We wish to consider the length L_n of the longest common subsequence (LCS) of two A-valued i.i.d. sequences X_1, \ldots, X_n and Y_1, \ldots, Y_n with a common law μ . This problem and its variants have been much studied in probability theory [6, 7, 19], computer science [1, 3, 14] and mathematical biology [11, 15, 16, 18]; see also the volume [17] for several articles. In mathematical biology, the alphabet $A = \{a, c, t, g\}$ of greatest interest consists of the four DNA bases, and one may want to test whether an observed common subsequence between two base sequences could be due to chance. The quantity $2(n - L_n)$ is the minimal number of insertions and deletions needed to change either sequence to the other one; in computer science this "edit distance" is used as a metric on strings.

It is easy to see that $\{EL_n, n \ge 1\}$ is a superadditive sequence, that is,

(1.1)
$$EL_{n+m} \ge EL_n + EL_m \text{ for all } n, m \ge 1.$$

It therefore follows from standard methods that EL_n/n has a limit and the convergence is from below, that is, there exists $\gamma = \gamma(\mu) \in [0, 1]$ such that

(1.2)
$$\lim_{n} EL_{n}/n = \sup_{n} EL_{n}/n = \gamma.$$

Kingman's subadditive ergodic theorem [13] further implies that $L_n/n \rightarrow \gamma$ a.s. For fair coin tossing, where $A = \{H, T\}$ and $\mu(H) = \mu(T) = 1/2$, simula-

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tions and heuristics (see [19] and Section 3 below) suggest that γ is between 0.81 and 0.82.

What interests us here is the rate at which the convergence in (1.2) occurs. The following is our main result.

THEOREM 1.1. There exists a constant C such that for every alphabet A, law μ and $n \geq 1$,

(1.3)
$$\gamma n \ge EL_n \ge \gamma n - C(n \log n)^{1/2}.$$

For a given n_0 our calculations will give an explicit value of C valid for $n \ge n_0$. This C will be smaller for larger n_0 as lower-order terms become negligible. In fact, we will show in Section 2 that for any C > 3.42, (1.3) is valid for all sufficiently large n.

The bound in (1.3) is useful in conjunction with simulations, which can really only estimate EL_n , in estimating γ . Simulations with n = 100,000 will be discussed in Section 3, together with simulations which suggest that the $(n \log n)^{1/2}$ rate in Theorem 1.1 is nearly the best obtainable by the methods of this paper.

Our method is modeled after that of [2], where a rate of convergence problem for first-passage percolation was considered. The analog of L_n is the passage time from the origin to a point n units out on an axis. The applicability of the method is not surprising in view of the fact that the LCS problem can be reformulated as a dependent first-passage percolation problem, as noted in [1, 4, 14, 20].

2. Proof of the main result. In place of L_n it is more convenient to work with

$$U_n \coloneqq 2(n - L_n).$$

If we think of the corresponding letters of the two maximal identical subsequences as being matched, then U_n represents the number of letters unused in this matching. More generally, for $1 \le i \le j + 1$ and $1 \le m \le n + 1$, we define U([i, j], [m, n]) to be the number of letters unused after matching the corresponding letters of a longest common subsequence of X_i, \ldots, X_j and Y_m, \ldots, Y_n . When j + 1 = i and/or n + 1 = m, we interpret the corresponding sequence here as being empty, so that, for example, U([i, i - 1], [m, n]) = n - m + 1 if $m \le n + 1$. When j + 1 < i and/or n + 1 < m, we use the convention that $U([i, j], [m, n]) = \infty$. We will abbreviate U([1, j], [1, n]) to U(j, n). Define

$$V_n := \min_{-n \le k \le n} U(n+k, n-k).$$

These quantities appear naturally in the first-passage reformulation of LCS, so we will briefly describe that reformulation now. Consider the integer lattice in $[0, 2n] \times [0, 2n]$, with horizontal and vertical bonds between nearest-neighbor sites of the lattice (that is, pairs x, y with |x - y| = 1) and a

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diagonal bond from each (i - 1, j - 1) to $(i, j), 1 \le i \le 2n$ and $1 \le j \le 2n$. The passage time of each horizontal and vertical bond is defined to be 1, and the passage time of the diagonal bond from (i - 1, j - 1) to (i, j) is 0 if $X_i = Y_j$, and ∞ otherwise. Then U([i, j], [r, s]) represents the minimal total passage time among all paths from (i - 1, r - 1) to (j, s) for which each coordinate in nondecreasing. We will call such a path a nondecreasing path. Let l_n denote the diagonal from (0, 2n) to (2n, 0). Then V_n represents the minimal total passage time among all nondecreasing paths starting at (0, 0)and ending on l_n .

For $n \ge 1$ and $\beta > 0$, define the generating functions

$$g_n(\beta) := -\log \bigg(\sum_{-n \le k \le n} E \exp(-\beta (U(n+k, n-k) - 2)) \bigg).$$

Heuristically one expects the sum in the definition of $g_n(\beta)$ to behave like its largest term, so that $g_n(\beta)$ behaves like $-\log E \exp(-\beta V_n)$. In fact, by Jensen's inequality we have

$$(2.1) g_n(\beta) \leq -\log(Ee^{-\beta(V_n-2)}) \leq \beta(EV_n-2).$$

The key property of $g_n(\beta)$ is given in the following result.

PROPOSITION 2.1. For each $\beta > 0$, the sequence $\{g_n(\beta): n \ge 1\}$ is superadditive, that is,

(2.2)
$$g_{n+m}(\beta) \ge g_n(\beta) + g_m(\beta) \quad \text{for all } m, n \ge 0.$$

Consequently, for some constants $\nu_{\beta} \leq 2(1 - \gamma)$,

(2.3)
$$\lim_{n} g_{n}(\beta)/n = \sup_{n} g_{n}(\beta)/n = \beta \nu_{\beta} \quad \text{for each } \beta > 0.$$

Before proving Proposition 2.1 we note that together with (1.2) it tells us that, for each fixed n and β , we have

(2.4)
$$g_n(\beta)/\beta \le \nu_\beta n \le 2(1-\gamma)n \le EU_n.$$

Thus because EU_n is subadditive and $g_n(\beta)$ is superadditive, $g_n(\beta)/\beta$ and EU_n are on opposite sides of the limiting approximation $2(1 - \gamma)n$. It follows that

(2.5)
$$\frac{EU_{2n} - 4(1 - \gamma)n \leq EU_{2n} - 2g_n(\beta)/\beta}{= (EU_{2n} - 2EV_n) + 2(EV_n - g_n(\beta)/\beta).}$$

Note that all of this is valid even if β is chosen depending on n.

To prove Proposition 2.1 we will need the following result.

LEMMA 2.2. For each
$$n, m \ge 0$$
 and $0 \le k \le n + m$,
 $U(n + m + k, n + m - k) + 2$
 $\ge \min\{U([1, n + j], [1, n - j]) + U([n + j + 1, n + m + k], [n - j + 1, n + m - k]): -n \le j \le n, k - m \le j \le k + m\}.$

PROOF. Let Γ be a nondecreasing path of minimal total passage time from (0,0) to (n + m + k, n + m - k). Then Γ intersects l_n in a unique point (n + x, n - x), and for some integer $j, -n \le j < n$, either x = j or x = j + 1/2. Since the path is nondecreasing, we have $-n \le x \le n$ and $k - m \le x \le k + m$. If x = j, then breaking Γ into two pieces at (n + j, n - j) shows that

$$U(n + m + k, n + m - k)$$

= $U([1, n + j], [1, n - j])$
+ $U([n + j + 1, n + m + k], [n - j + 1, n + m + k]).$

If x = j + 1/2, then replacing the bond from (n + j, n - j - 1) to (n + j + 1, n - j) in Γ with the bond from (n + j, n - j - 1) to (n + j, n - j) and the bond from (n + j, n - j) to (n + j + 1, n - j) adds 2 to the passage time. Breaking the altered Γ at (n + j, n - j) then shows that

$$U(n + m + k, n + m - k) + 2$$

$$\geq U([1, n + j], [1, n - j])$$

$$+ U([n + j + 1, n + m + k], [n - j + 1, n + m - k]).$$

In both cases, the desired result follows. \Box

PROOF OF PROPOSITION 2.1. From Lemma 2.2, independence and translation invariance,

$$\begin{split} \sum_{-(n+m) \le k \le n+m} & E \exp \left(-\beta \left[U(n+m+k,n+m-k) - 2 \right] \right) \\ \le \sum_{-(n+m) \le k \le n+m} & E \exp \left(-\beta \min \{ U([1,n+j],[1,n-j]) \\ & + U([n+j+1,n+m+k],[n-j+1,n+m-k]) - 4; \\ & -n \le j \le n, k-m \le j \le k+m \} \right) \\ \le \sum_{-(n+m) \le k \le n+m} & \sum_{\substack{j: -n \le j \le n \\ k-m \le j \le k+m}} & E \exp \left(-\beta \{ U([1,n+j],[1,n-j]) \\ & + U([n+j+1,n+m+k],[n-j+1,n+m-k]) - 4 \} \right) \\ = & \sum_{-(n+m) \le k \le n+m} & \sum_{\substack{j: -n \le j \le n \\ k-m \le j \le k+m}} & E \exp \left(-\beta [U(n+j,n-j) - 2] \right) \\ & = & \left(\sum_{j: -n \le j \le n} & E \exp \left(-\beta [U(n+r,m-r) - 2] \right) \right) \\ & \times \left(\sum_{r: -m \le r \le m} & E \exp \left(\{ \beta [U(m+r,m-r) - 2]) \right) \right). \end{split}$$

Taking logarithms then yields (2.2). Standard subadditivity arguments then give (2.3). The fact that $\nu_{\beta} \leq 2(1 - \gamma)$ follows from (2.1) and the fact that $V_n \leq U_n$. \Box

The following lemma gives a special case of Azuma's inequality [5] and is essentially a martingale version of Theorem 2 of Hoeffding [10].

LEMMA 2.3. Suppose $f(x_1, \ldots, x_n, y_1, \ldots, y_n)$ is a function on A^{2n} with the property that changing any one argument of f while holding the others fixed changes the value of f by at most 2. Then for $Z := f(X_1, \ldots, X_n, Y_1, \ldots, Y_n)$ and $u \ge 0$,

$$P[Z - EZ \ge u] \le \exp(-u^2/4n).$$

In particular,

(2.6)
$$P[L_{2n} - EL_{2n} \ge u/2] = P[U_{2n} - EU_{2n} \le -u] \le \exp(-u^2/8n),$$

(2.7)
$$P[L_{2n} - EL_{2n} \le -u/2] = P[U_{2n} - EU_{2n} \ge u] \le \exp(-u^2/8n),$$

(2.8)
$$P[V_n - EV_n \le -u] \le \exp(-u^2/8n)$$

and

(2.9)
$$P[V_n - EV_n \ge u] \le \exp(-u^2/8n).$$

Theorem 1.1 is an immediate consequence of the next proposition, since monotonicity of EU_n handles odd indices n. Any fixed values of $\lambda > 1$ and $\theta > \sqrt{2}$ suffice for proving Theorem 1.1, so if $C > (2 + \sqrt{2})$, then (1.3) is valid for all sufficiently large n. For the explicit confidence intervals of Section 3, the more detailed requirements in (2.10) and (2.11) are important.

PROPOSITION 2.4. Suppose $n \ge 8$ and $\lambda, \theta > 0$ satisfy

(2.10)
$$\lambda^2 \ge 1 + 1/(2n\log 2n) + 2\lambda/(2n\log 2n)^{1/2}$$

 $+(\log 5.1\lambda)/\log 2n + (\log \log 2n)/(2\log 2n)$

and

(2.11)
$$\theta^2 \ge 2 + (\log 4) / \log 2n.$$

Then

(2.12)
$$\frac{EU_{2n}/(2n) \le 2(1-\gamma) + 2(2\lambda+\theta)((\log 2n)/(2n))^{1/2}}{+((8\log 2)/n)^{1/2}}$$

and

(2.13)

$$EL_{2n}/(2n) \ge \gamma - (2\lambda + \theta) ((\log 2n)/(2n))^{1/2} - ((2\log 2)/n)^{1/2}.$$

PROOF. We will use (2.5). Let

$$\beta_n \coloneqq \lambda((\log 2n)/(2n))^{1/2}.$$

Let us first bound $EV_n - g_n(\beta_n)/\beta_n$. From integration by parts and Lemma 2.3,

$$\begin{split} E \exp(-\beta_n V_n) \\ &= \int_0^\infty \beta_n \exp(-\beta_n x) P[V_n \le x] dx \\ &\le \exp(-\beta_n E V_n) + \int_0^{E V_n} \beta_n \exp(-\beta_n x) \exp(-(E V_n - x)^2 / (8n)) dx \\ &\le \exp(-\beta_n E V_n) + \beta_n (8\pi n)^{1/2} \exp(-\beta_n E V_n + 2n\beta_n^2). \end{split}$$

Therefore,

$$\begin{split} \exp(-g_n(\beta_n)) &= \sum_{-n \le k \le n} E \exp(-\beta_n(U(n+k,n-k)-2)) \\ (2.14) &\leq (2n+1) E \exp(-\beta_n(V_n-2)) \\ &\leq (2n+1) \{\exp(-\beta_n(EV_n-2))\} \Big[1 + \beta_n(8\pi n)^{1/2} \exp(2n\beta_n^2) \Big]. \end{split}$$

Observe that

(2.15)
$$\begin{aligned} \log\left(1.011 \left(8\pi n\,\beta_n^2\right)^{1/2} \exp(2n\,\beta_n^2)\right) \\ \leq \log(2.860\pi^{1/2}\lambda) + \log\left(\left(n\,\beta_n^2\right)^{1/2}/\lambda\right) + 2n\,\beta_n^2 \end{aligned}$$

Taking logs in (2.14), rearranging and using $n\beta_n^2 \ge 2\log 2$, (2.15) and (2.10), we obtain

$$\begin{split} EV_n &- g_n(\beta_n)/\beta_n \\ &\leq 2 + \beta_n^{-1} \log(2n+1) + \beta_n^{-1} \log\left(1 + \left(8\pi n\beta_n^2\right)^{1/2} \exp(2n\beta_n^2)\right) \\ &\leq 2 + \beta_n^{-1} (\log 2n + 1/(2n)) + \beta_n^{-1} \log\left(1.011 \left(8\pi n\beta_n^2\right)^{1/2} \exp(2n\beta_n^2)\right) \\ (2.16) &\leq \lambda^{-1} (2n\log 2n)^{1/2} \\ &\times \left[1 + 1/(2n\log 2n) + 2\lambda/(2n\log 2n)^{1/2} \\ &+ (\log 5.1\lambda)/\log 2n + (\log \log 2n)/2\log 2n + \lambda^2\right] \end{split}$$

 $\leq 2\lambda \big(2n\log 2n\big)^{1/2}.$

Let us next bound $EU_{2n} - 2EV_n$. In terms of the first-passage formulation, we use what is essentially a reflection argument across the diagonal l_n . From (2.6) of Lemma 2.3, we have

(2.17)
$$1/2 \le P \Big[V_n \le EV_n + (8n \log 2)^{1/2} \Big] \le \sum_{-n < j < n} P \Big[U(n+j, n-j) \le EV_n + (8n \log 2)^{1/2} \Big],$$

where j = n and -n need not be included in the sum because the minimum of U(n + j, n - j) always occurs with -n < j < n. Therefore, there exists an

index j with

$$P\Big[U(n+j, n-j) \le EV_n + (8n\log 2)^{1/2}\Big] \ge 1/(4n-2).$$

Since U([n + j + 1, 2n], [n - j + 1, 2n]) has the same distribution as U(n + j, n - j) = U([1, n + j], [1, n - j]) and is independent of it, we have

$$1/(4n-2)^{2} \leq P[U([1, n+j], [1, n-j])] \leq EV_{n} + (8n \log 2)^{1/2},$$

(2.18) $U([n+j+1,2n],[n-j+1,2n]) \le EV_n + (8n\log 2)^{1/2}]$

$$\leq P \Big[U_{2n} \leq 2 \Big(EV_n + (8n \log 2)^{1/2} \Big) \Big].$$

However, from (2.6) of Lemma 2.3 and (2.11), (2.19) $P\left[U_{2n} \le EU_{2n} - 2\theta(2n\log 2n)^{1/2}\right] \le \exp(-\theta^2 \log 2n) \le 1/(4n)^2$, which with (2.18) shows

$$EU_{2n} \le 2EV_n + 2(8n\log 2)^{1/2} + 2\theta(2n\log 2n)^{1/2}.$$

With (2.5) and (2.16) this proves (2.12), which implies (2.13). \Box

Our method should be applicable to other problems which have a firstpassage formulation. The main ingredients for which one must have analogs are, first, that there is enough independence that something like Lemma 2.3 holds, and second, that for a path Γ as in the proof of Lemma 2.2 which meets l_n at a particular point, the two segments into which Γ is split by that point are independent or at least are appropriately comparable to independent segments as in [2].

3. Confidence bounds and simulations. For $2n \ge 100,000$, (2.10) and (2.11) are satisfied with $\lambda = 1.123$ and $\theta = 1.457$. Therefore, from Proposition 2.4,

(3.1)
$$EL_{2n}/2n \le \gamma \le EL_{2n}/(2n) + 0.0450.$$

Given k independent observations of L_{2n} , let \overline{L}_{2n} denote the sample mean of these observations. For optimal confidence bounds, we place our estimate of γ in the center of the interval suggested by (3.1), defining

(3.2)
$$\hat{\gamma}_{2n} \coloneqq \overline{L}_{2n}/(2n) + 0.0225.$$

From Lemma 2.3, for $2n \ge 100,000$, $x \ge 0.0055/\sqrt{k}$ and $t := \gamma - EL_{2n}/2n$, $P[|\hat{\gamma}_{0n} - \gamma| > x + 0.0225]$

$$\leq P \Big[\overline{L}_{2n} / (2n) - EL_{2n} / (2n) \ge x + t \Big] \\ + P \Big[\overline{L}_{2n} / (2n) - EL_{2n} / (2n) \le -(x + 0.0450 - t) \Big] \\ \leq \exp \Big(-2kn(x + t)^2 \Big) 2 + \exp \Big(-2kn(x + 0.0450 - t)^2 \Big) \\ \leq \exp (-2knx^2) + \exp \Big(-2kn(x + 0.0450)^2 \Big) \\ \le 0.05.$$

In particular, for $2n \ge 100,000$ and k = 2,

(3.3)
$$P[|\hat{\gamma}_{2n} - \gamma| \le 0.0264] \ge 0.95$$

Eggert and Waterman [9] simulated two trials of L_{2n} for fair coin tossing with 2n = 100,000. The observed values were 81,223 and 81,146, yielding the estimate $\hat{\gamma}_{2n} = 0.8343$. Therefore, from (3.3),

$$(3.4) 0.8079 \le \gamma \le 0.8607.$$

with 95% confidence. By contrast, the best bounds known with certainty for fair coin tossing are $0.7615 \le \gamma \le 0.8376$ [7, 8]. By using this upper bound we can improve on (3.4) as follows. From Lemma 2.3 we have $P[\bar{L}_{2n}/2n - \delta \le \gamma] \ge 1 - \exp(-2kn\delta^2)$. In particular, with k = 2 and $\delta = 0.0039$ this yields $0.8079 \le \gamma \le 0.8376$ with 95% confidence.

Additional simulations in [21] suggest that the variance of L_n is approximately proportional to n. If this is true, then Lemma 2.3 cannot be valid with any smaller power of n in the denominator of the exponent. This means that our method cannot yield a better power of n than the 1/2 which appears in (1.3). Of course, the actual difference $\gamma n - EL_n$ may well be $o(n^{1/2})$, but quite different methods would apparently be needed to obtain such an improved result.

NOTE ADDED IN PROOF. An additional reference by P. Jaillet [(1992) Math Oper. Res. 17 964–980), has a proof of Lemma 2.3 (ii) and gives two-sided bounds with rates as in (1.4) for several functionals including TSP and MST.

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